



LAMPIRAN

LAMPIRAN 1

Bukti Lemma 2.18.

Misalkan matriks A berukuran $n \times n$ merupakan matriks definit positif. Lemma 2.21 akan dibuktikan dengan menggunakan kontradiksi. Andaikan A merupakan matriks yang *singular* atau secara ekivalen, yaitu $\text{rank}(A) < n$. Akibatnya, kolom A tidak bebas linier sehingga terdapat vektor tidak nol x_* , sedemikian sehingga $A x_* = \mathbf{0}$. Oleh karena itu,

$$x_*^T A x_* = x_*^T (A x_*) = x_*^T \mathbf{0} = 0.$$

Padahal A merupakan matriks definit positif, artinya

$$x^T A x > 0, \text{ untuk setiap } x \neq 0.$$

Sedangkan, jika A merupakan matriks *singular* diperoleh bahwa:

$$x_*^T A x_* = x_*^T (A x_*) = x_*^T \mathbf{0} = 0$$

Berarti kontradiksi dengan pengandaian yang digunakan sehingga sembarang matriks definit positif adalah *nonsingular*.

LAMPIRAN 2

Bukti (2.3.2.a), (2.3.2.b), dan (2.3.2.c).

Misalkan a_i dinyatakan sebagai entri ke- i dari \mathbf{a} dan x_i sebagai entri ke- i dari \mathbf{x} , sedangkan a_{ik} dinyatakan sebagai entri ke- ik dari \mathbf{A} . Maka

$$\mathbf{a}^T \mathbf{x} = \sum_i a_i x_i \text{ dan } \mathbf{x}^T \mathbf{A} \mathbf{x} = \sum_{i,k} a_{ik} x_i x_k$$

Sebagai langkah awal akan didefinisikan:

$$\frac{\partial x_i}{\partial x_j} = \begin{cases} 1, & \text{jika } i = j \\ 0, & \text{jika } i \neq j \end{cases}$$

$$\frac{\partial(x_i, x_k)}{\partial x_j} = \begin{cases} 2x_j, & \text{jika } i = k = j \\ x_i, & \text{jika } k = j \text{ tetapi } i \neq j \\ x_k, & \text{jika } i = j \text{ tetapi } k \neq j \\ 0, & \text{lainnya.} \end{cases}$$

sehingga diperoleh

$$\frac{\partial(\mathbf{a}^T \mathbf{x})}{\partial x_j} = \frac{\partial(\sum_i a_i x_i)}{\partial x_j} = \sum_i a_i \frac{\partial x_i}{\partial x_j} = a_j \quad (\text{L2.1})$$

dan

$$\frac{\partial(\mathbf{x}' \mathbf{A} \mathbf{x})}{\partial x_j} = \frac{\partial(\sum_{i,k} a_{ik} x_i x_k)}{\partial x_j}$$

$$\begin{aligned}
&= \frac{\partial(a_{jj}x_j^2 + \sum_{i \neq j} a_{ij}x_i x_j + \sum_{k \neq j} a_{jk}x_j x_k + \sum_{i \neq j, k \neq j} a_{ik}x_i x_k)}{\partial x_j} \\
&= a_{jj} \frac{\partial(x_j^2)}{\partial x_j} + \sum_{i \neq j} a_{ij} \frac{\partial(x_i x_j)}{\partial x_j} + \sum_{k \neq j} a_{jk} \frac{\partial(x_j x_k)}{\partial x_j} + \sum_{i \neq j, k \neq j} a_{ik} \frac{\partial(x_i x_k)}{\partial x_j} \\
&= 2a_{jj}x_j + \sum_{i \neq j} a_{ij}x_i + \sum_{k \neq j} a_{jk}x_k + 0 \\
&= \sum_i a_{ij}x_i + \sum_k a_{jk}x_k. \tag{L2.2}
\end{aligned}$$

Persamaan (L2.1) dapat dibentuk dalam notasi vektor sebagai

$$\frac{\partial(\mathbf{a}^T \mathbf{x})}{\partial \mathbf{x}} = \mathbf{a} \quad \text{atau} \quad \frac{\partial(\mathbf{a}^T \mathbf{x})}{\partial \mathbf{x}^T} = \mathbf{a}^T$$

Sedangkan persamaan (L2.2) dapat dibentuk dalam notasi matriks sebagai:

$$\frac{\partial(\mathbf{x}^T \mathbf{A} \mathbf{x})}{\partial \mathbf{x}} = (\mathbf{A} + \mathbf{A}^T)\mathbf{x}$$

dengan $\sum_k a_{jk}x_k$ adalah entri ke- j dari vektor kolom $\mathbf{A}\mathbf{x}$ dan $\sum_i a_{ij}x_i$ adalah entri ke- j dari $\mathbf{A}^T\mathbf{x}$.

LAMPIRAN 3

Bukti (2.3.3.a).

Misalkan $\mathbf{L} = a\mathbf{F} + b\mathbf{G}$. Elemen ke- i s dari \mathbf{L} adalah sebagai berikut:

$$l_{is} = af_{is} + bg_{is}.$$

Fungsi f_{is} dan g_{is} *continuously differentiable* di \mathbf{c} sehingga l_{is} juga *continuously differentiable* di $\mathbf{x} = \mathbf{c}$, yakni:

$$\frac{\partial l_{is}}{\partial x_j} = a \frac{\partial f_{is}}{\partial x_j} + b \frac{\partial g_{is}}{\partial x_j}.$$

Oleh karena $\partial l_{is}/\partial x_j$, $\partial f_{is}/\partial x_j$, dan $\partial g_{is}/\partial x_j$ masing-masing merupakan elemen ke- i s dari $\partial \mathbf{L}/\partial x_j$, $\partial \mathbf{F}/\partial x_j$, dan $\partial \mathbf{G}/\partial x_j$, berarti \mathbf{L} juga *continuously differentiable* di $\mathbf{x} = \mathbf{c}$, yaitu:

$$\frac{\partial \mathbf{L}}{\partial x_j} = a \frac{\partial \mathbf{F}}{\partial x_j} + b \frac{\partial \mathbf{G}}{\partial x_j}.$$

LAMPIRAN 4

Bukti (2.3.3.b).

Misalkan $\mathbf{H} = \mathbf{FG}$. Elemen ke- it dari \mathbf{H} adalah sebagai berikut:

$$h_{it} = \sum_{s=1}^q f_{is} g_{st} .$$

Fungsi f_{is} dan g_{st} continuously differentiable di \mathbf{c} sehingga $f_{is} g_{st}$ juga continuously differentiable di $\mathbf{x} = \mathbf{c}$, yakni:

$$\frac{\partial f_{is} g_{st}}{\partial x_j} = f_{is} \frac{\partial g_{st}}{\partial x_j} + \frac{\partial f_{is}}{\partial x_j} g_{st} .$$

Oleh karena itu, h_{it} continuously differentiable di $\mathbf{x} = \mathbf{c}$ atau dapat dituliskan

$$\frac{\partial h_{it}}{\partial x_j} = \sum_{s=1}^q \frac{\partial (f_{is} g_{st})}{\partial x_j} = \sum_{s=1}^q f_{is} \frac{\partial g_{st}}{\partial x_j} + \sum_{s=1}^q \frac{\partial f_{is}}{\partial x_j} g_{st} .$$

Oleh karena $\sum_{s=1}^q f_{is} (\partial g_{st} / \partial x_j)$ dan $\sum_{s=1}^q (\partial f_{is} / \partial x_j) g_{st}$ masing-masing adalah elemen ke- it dari $\mathbf{F}(\partial \mathbf{G} / \partial x_j)$ dan $(\partial \mathbf{F} / \partial x_j) \mathbf{G}$, dapat disimpulkan bahwa \mathbf{H} juga continuously differentiable di $\mathbf{x} = \mathbf{c}$, yaitu:

$$\frac{\partial \mathbf{FG}}{\partial x_j} = \mathbf{F} \frac{\partial \mathbf{G}}{\partial x_j} + \frac{\partial \mathbf{F}}{\partial x_j} \mathbf{G} .$$

LAMPIRAN 5

Bukti (2.3.7.a), (2.3.7.b), dan (2.3.7.c).

Berdasarkan (2.3.3.b), (2.3.4.a), dan (2.3.6.c) didapatkan bahwa

$$\begin{aligned}
 \frac{\partial^2 \det(\mathbf{F})}{\partial x_i \partial x_j} &= \frac{\partial \left\{ |\mathbf{F}| \operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \right) \right\}}{\partial x_i} \\
 &= |\mathbf{F}| \left[\operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i \partial x_j} \right) + \operatorname{tr} \left(-\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \right) \right] \\
 &\quad + |\mathbf{F}| \operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \right) \operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \right) \\
 &= |\mathbf{F}| \left[\operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i \partial x_j} \right) + \operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \right) \operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \right) \right] \\
 &\quad - \operatorname{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \right)
 \end{aligned}$$

dan dengan menggunakan (2.3.3.c) dan (2.3.6.e) didapatkan pula bahwa

$$\begin{aligned}
 \frac{\partial^2 \mathbf{F}^{-1}}{\partial x_i \partial x_j} &= \frac{\partial \left[-\mathbf{F}^{-1} \left(\frac{\partial \mathbf{F}}{\partial x_j} \right) \mathbf{F}^{-1} \right]}{\partial x_i} \\
 &= -\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \frac{\partial \mathbf{F}^{-1}}{\partial x_i} - \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i \partial x_j} \mathbf{F}^{-1} - \frac{\partial \mathbf{F}^{-1}}{\partial x_i} \frac{\partial \mathbf{F}}{\partial x_j} \mathbf{F}^{-1} \\
 &= -\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \left(-\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1} \right) - \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i \partial x_j} \mathbf{F}^{-1} \\
 &\quad - \left(-\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1} \right) \frac{\partial \mathbf{F}}{\partial x_j} \mathbf{F}^{-1}
 \end{aligned}$$

$$\begin{aligned}
&= -\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i \partial x_j} \mathbf{F}^{-1} + \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \mathbf{F}^{-1} \\
&\quad + \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1}.
\end{aligned}$$

Selanjutnya, berdasarkan (2.3.4.a), (2.3.6.d), dan (2.3.6.e) didapatkan bahwa

$$\begin{aligned}
\frac{\partial^2 \log \det(\mathbf{F})}{\partial x_i \partial x_j} &= \frac{\partial \text{tr} \left[\mathbf{F}^{-1} \left(\frac{\partial \mathbf{F}}{\partial x_j} \right) \right]}{\partial x_i} \\
&= \text{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i \partial x_j} \right) + \text{tr} \left(-\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \right) \\
&= \text{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i \partial x_j} \right) - \text{tr} \left(\mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_i} \mathbf{F}^{-1} \frac{\partial \mathbf{F}}{\partial x_j} \right).
\end{aligned}$$

LAMPIRAN 6

Bukti θ memaksimumkan $L(\theta) \leftrightarrow \theta$ memaksimumkan $\ln L(\theta)$.

Keterangan: $\theta = (\theta_1, \theta_2, \dots, \theta_n)$.

Bukti:

(\rightarrow) Oleh karena θ memaksimumkan $L(\theta)$, maka:

- $\frac{\partial L(\theta)}{\partial \theta_1} = 0,$
- $\frac{\partial L(\theta)}{\partial \theta_2} = 0,$
⋮
- $\frac{\partial L(\theta)}{\partial \theta_n} = 0,$
- $\frac{\partial^2 L(\theta)}{\partial \theta_i^2} < 0, \quad i = 1, 2, \dots, n,$
- $D = \frac{\partial^2 L(\theta)}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\theta)}{\partial \theta_j^2} - \left(\frac{\partial^2 L(\theta)}{\partial \theta_i \partial \theta_j} \right)^2 > 0, \quad i = 1, 2, \dots, n, \quad j = 1, 2, \dots, n, \quad i \neq j.$

Akan ditunjukkan bahwa θ juga memaksimumkan $\ln L(\theta)$, yaitu:

- $\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_1} = 0,$

$$\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_1} = \frac{1}{L(\boldsymbol{\theta})} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_1} = \frac{1}{L(\boldsymbol{\theta})} \cdot 0 = 0$$

$$\therefore \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_1} = 0,$$

- $\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_2} = 0$

$$\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_2} = \frac{1}{L(\boldsymbol{\theta})} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_2} = \frac{1}{L(\boldsymbol{\theta})} \cdot 0 = 0$$

$$\therefore \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_2} = 0,$$

⋮

- $\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_n} = 0$

$$\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_n} = \frac{1}{L(\boldsymbol{\theta})} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_n} = \frac{1}{L(\boldsymbol{\theta})} \cdot 0 = 0$$

$$\therefore \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_n} = 0,$$

- $\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} < 0, \quad i = 1, 2, \dots, n,$

$$\begin{aligned}
\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} &= \frac{\partial}{\partial \theta_i} \left(\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_i} \right) = \frac{\partial}{\partial \theta_i} \left(\frac{1}{L(\boldsymbol{\theta})} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right) \\
&= \frac{\partial}{\partial \theta_i} \left(\frac{1}{L(\boldsymbol{\theta})} \right) \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} + \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{1}{L(\boldsymbol{\theta})} = \frac{\partial}{\partial \theta_i} \left(\frac{1}{L(\boldsymbol{\theta})} \right) \cdot 0 + \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{1}{L(\boldsymbol{\theta})} \\
&= \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{1}{L(\boldsymbol{\theta})} < 0 \cdot \frac{1}{L(\boldsymbol{\theta})} = 0
\end{aligned}$$

$$\therefore \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} < 0, \quad i = 1, 2, \dots, n,$$

- $D = \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 > 0, \quad i = 1, 2, \dots, n, j = 1, 2, \dots, n, i \neq j$

$$\begin{aligned}
\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} &= \frac{\partial}{\partial \theta_i} \left(\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_i} \right) = \frac{\partial}{\partial \theta_i} \left(\frac{\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i}}{L(\boldsymbol{\theta})} \right), \quad i = 1, 2, \dots, n \\
&= \frac{\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot L(\boldsymbol{\theta}) - \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i}}{L^2(\boldsymbol{\theta})}
\end{aligned}$$

$$\begin{aligned}
\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} &= \frac{\partial}{\partial \theta_j} \left(\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_j} \right) = \frac{\partial}{\partial \theta_j} \left(\frac{\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j}}{L(\boldsymbol{\theta})} \right), \quad j = 1, 2, \dots, n \\
&= \frac{\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} \cdot L(\boldsymbol{\theta}) - \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j}}{L^2(\boldsymbol{\theta})}
\end{aligned}$$

$$\begin{aligned} \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} &= \frac{\partial}{\partial \theta_i} \left(\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_j} \right) = \frac{\partial}{\partial \theta_i} \left(\frac{\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j}}{L(\boldsymbol{\theta})} \right), i = 1, 2, \dots, n, j = 1, 2, \dots, n \\ &= \frac{\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \cdot L(\boldsymbol{\theta}) - \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j}}{L^2(\boldsymbol{\theta})}. \end{aligned}$$

$$\begin{aligned} D &= \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \\ &= \left[\frac{\left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot L(\boldsymbol{\theta}) - \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)}{L^2(\boldsymbol{\theta})} \right] \cdot \left[\frac{\left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} \cdot L(\boldsymbol{\theta}) - \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)}{L^2(\boldsymbol{\theta})} \right] \\ &\quad - \left[\frac{\left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \cdot L(\boldsymbol{\theta}) - \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)}{L^2(\boldsymbol{\theta})} \right]^2, i = 1, 2, \dots, n, j = 1, 2, \dots, n \\ &= \frac{1}{L^4(\boldsymbol{\theta})} \left[\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} \cdot L^2(\boldsymbol{\theta}) + \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 \right. \\ &\quad \left. - \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 \cdot L(\boldsymbol{\theta}) - \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} \cdot \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 \cdot L(\boldsymbol{\theta}) \right] \\ &\quad - \frac{1}{L^4(\boldsymbol{\theta})} \left[\left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \cdot L^2(\boldsymbol{\theta}) + \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 \right. \\ &\quad \left. - 2 \left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \cdot L(\boldsymbol{\theta}) \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right) \right] \end{aligned}$$

$$\begin{aligned}
&= \frac{L^2(\boldsymbol{\theta})}{L^4(\boldsymbol{\theta})} \left[\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \right] \\
&\quad + \frac{L(\boldsymbol{\theta})}{L^4(\boldsymbol{\theta})} \left[2 \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} - \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 \right. \\
&\quad \left. - \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} \cdot \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 \right] \\
&= \frac{1}{L^2(\boldsymbol{\theta})} \left[\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \right] \\
&\quad + \frac{1}{L^3(\boldsymbol{\theta})} \left[2 \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \cdot 0 \cdot 0 - \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot (0)^2 - \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} \cdot (0)^2 \right] \\
&= \frac{1}{L^2(\boldsymbol{\theta})} \left[\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \right] > \frac{1}{L^2(\boldsymbol{\theta})} \cdot 0 = 0 \\
\therefore D &= \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 > 0, \quad i = 1, 2, \dots, n, j = 1, 2, \dots, n, i \neq j.
\end{aligned}$$

Dengan perkataan lain terbukti bahwa $\boldsymbol{\theta}$ juga memaksimumkan $\ln L(\boldsymbol{\theta})$.

(\leftarrow) Oleh karena $\boldsymbol{\theta}$ memaksimumkan $\ln L(\boldsymbol{\theta})$, maka

- $\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_1} = 0,$

- $\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_2} = 0,$

\vdots

- $\frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_n} = 0,$
- $\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} < 0, \quad i = 1, 2, \dots, n,$
- $D = \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 > 0, \quad i = 1, 2, \dots, n, j = 1, 2, \dots, n, i \neq j.$

Akan ditunjukkan bahwa $\boldsymbol{\theta}$ juga memaksimumkan $L(\boldsymbol{\theta})$, yaitu:

$$\bullet \quad \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_1} = 0,$$

$$\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_1} = L(\boldsymbol{\theta}) \cdot \frac{1}{L(\boldsymbol{\theta})} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_1} = L(\boldsymbol{\theta}) \cdot \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_1} = L(\boldsymbol{\theta}) \cdot 0 = 0$$

$$\therefore \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_1} = 0,$$

$$\bullet \quad \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_2} = 0,$$

$$\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_2} = L(\boldsymbol{\theta}) \cdot \frac{1}{L(\boldsymbol{\theta})} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_2} = L(\boldsymbol{\theta}) \cdot \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_2} = L(\boldsymbol{\theta}) \cdot 0 = 0$$

$$\therefore \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_2} = 0,$$

\vdots

- $\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_n} = 0,$

$$\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_n} = L(\boldsymbol{\theta}) \cdot \frac{1}{L(\boldsymbol{\theta})} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_n} = L(\boldsymbol{\theta}) \cdot \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_n} = L(\boldsymbol{\theta}) \cdot 0 = 0$$

$$\therefore \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_n} = 0,$$

- $\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} < 0, \quad i = 1, 2, \dots, n,$

$$\begin{aligned}\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} &= \frac{\partial \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)}{\partial \theta_i} \\ &= \frac{\partial}{\partial \theta_i} \left(L(\boldsymbol{\theta}) \cdot \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_i} \right) \\ &= \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial \ln L(\boldsymbol{\theta})}{\partial \theta_i} + L(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \\ &= \frac{1}{L(\boldsymbol{\theta})} \cdot \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 + L(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \\ &= \frac{1}{L(\boldsymbol{\theta})} \cdot \left[\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \right] \\ &= \frac{1}{L(\boldsymbol{\theta})} \cdot \left[(0)^2 + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \right] \\ &= L(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} < L(\boldsymbol{\theta}) \cdot 0 = 0\end{aligned}$$

$$\therefore \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} < 0, \quad i = 1, 2, \dots, n,$$

- $D = \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 > 0, \quad i = 1, 2, \dots, n, j = 1, 2, \dots, n, i \neq j$

$$\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} = \frac{\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2}}{L(\boldsymbol{\theta})}, \quad i = 1, 2, \dots, n$$

$$\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} = \frac{\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2}}{L(\boldsymbol{\theta})}, \quad j = 1, 2, \dots, n$$

$$\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} = \frac{\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j}}{L(\boldsymbol{\theta})}, \quad i = 1, 2, \dots, n, j = 1, 2, \dots, n$$

$$\begin{aligned} D &= \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \\ &= \left(\frac{\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2}}{L(\boldsymbol{\theta})} \right) \cdot \left(\frac{\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2}}{L(\boldsymbol{\theta})} \right) \\ &\quad - \left(\frac{\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} + L^2(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j}}{L(\boldsymbol{\theta})} \right)^2, \quad i = 1, 2, \dots, n, j = 1, 2, \dots, n, i \neq j \end{aligned}$$

$$\begin{aligned}
&= \frac{1}{L^2(\boldsymbol{\theta})} \cdot \left[\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 + L^4(\boldsymbol{\theta}) \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} + \right. \\
&\quad \left. + L^2(\boldsymbol{\theta}) \cdot \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} + L^2(\boldsymbol{\theta}) \cdot \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \right] \\
&- \frac{1}{L^2(\boldsymbol{\theta})} \cdot \left[\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 + L^4(\boldsymbol{\theta}) \cdot \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 + 2 \cdot L^2(\boldsymbol{\theta}) \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right] \\
&= \frac{L^4(\boldsymbol{\theta})}{L^2(\boldsymbol{\theta})} \cdot \left[\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \right] \\
&+ \frac{L^2(\boldsymbol{\theta})}{L^2(\boldsymbol{\theta})} \cdot \left[\left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \right)^2 \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} + \left(\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \right)^2 \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \right. \\
&\quad \left. - 2 \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_i} \cdot \frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right] \\
&= L^2(\boldsymbol{\theta}) \cdot \left[\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \right] \\
&+ 1 \cdot \left[(0)^2 \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} + (0)^2 \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} - 2 \cdot 0 \cdot 0 \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right] \\
&= L^2(\boldsymbol{\theta}) \cdot \left[\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 \ln L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 \right] > L^2(\boldsymbol{\theta}) \cdot 0 = 0
\end{aligned}$$

$\therefore D = \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i^2} \cdot \frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_j^2} - \left(\frac{\partial^2 L(\boldsymbol{\theta})}{\partial \theta_i \partial \theta_j} \right)^2 > 0, \quad i = 1, 2, \dots, n, j = 1, 2, \dots, n, i \neq j$

Dengan perkataan lain, terbukti bahwa $\boldsymbol{\theta}$ juga memaksimumkan $L(\boldsymbol{\theta})$.

Terbukti bahwa $\boldsymbol{\theta}$ memaksimumkan $L(\boldsymbol{\theta}) \leftrightarrow \boldsymbol{\theta}$ memaksimumkan $\ln L(\boldsymbol{\theta})$,

$$\boldsymbol{\theta} = (\theta_1, \theta_2, \dots, \theta_n).$$

LAMPIRAN 7

Metode Newton-Raphson.

Misalkan δ adalah vektor dari q -buah parameter yang tidak diketahui, yaitu

$\delta = (\delta_0, \delta_1, \dots, \delta_q)^T$, $l(\delta)$ adalah fungsi log-likelihood dari vektor δ . $s(\delta)$

adalah vektor turunan parsial pertama dari $l(\delta)$ dengan elemen

$$s_j(\delta) = \frac{\partial l(\delta)}{\partial \delta_j}, j = 0, 1, \dots, q, \text{ yaitu } s(\delta) = \begin{pmatrix} s_0(\delta) \\ s_1(\delta) \\ \vdots \\ s_q(\delta) \end{pmatrix} = \begin{pmatrix} \frac{\partial l(\delta)}{\partial \delta_0} \\ \frac{\partial l(\delta)}{\partial \delta_1} \\ \vdots \\ \frac{\partial l(\delta)}{\partial \delta_q} \end{pmatrix}, \text{ sedangkan}$$

$H(\delta)$ menyatakan matriks turunan parsial kedua dari $l(\delta)$ dengan elemen

$$H_{jk}(\delta) = \frac{\partial^2 l(\delta)}{\partial \delta_k \partial \delta_j}, j, k = 0, 1, \dots, q.$$

$$\begin{aligned} H(\delta) &= \frac{\partial s(\delta)}{\partial \delta^T} = \left(\begin{array}{cccc} \frac{\partial s(\delta)}{\partial \delta_0} & \frac{\partial s(\delta)}{\partial \delta_1} & \dots & \frac{\partial s(\delta)}{\partial \delta_q} \end{array} \right) \\ &= \left(\begin{array}{cccc} \frac{\partial s_0(\delta)}{\partial \delta_0} & \frac{\partial s_0(\delta)}{\partial \delta_1} & \dots & \frac{\partial s_0(\delta)}{\partial \delta_q} \\ \frac{\partial s_1(\delta)}{\partial \delta_0} & \frac{\partial s_1(\delta)}{\partial \delta_1} & \dots & \frac{\partial s_1(\delta)}{\partial \delta_q} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial s_q(\delta)}{\partial \delta_0} & \frac{\partial s_q(\delta)}{\partial \delta_1} & \dots & \frac{\partial s_q(\delta)}{\partial \delta_q} \end{array} \right) = \left(\begin{array}{cccc} \frac{\partial^2 l(\delta)}{\partial \delta_0^2} & \frac{\partial^2 l(\delta)}{\partial \delta_1 \partial \delta_0} & \dots & \frac{\partial^2 l(\delta)}{\partial \delta_q \partial \delta_0} \\ \frac{\partial^2 l(\delta)}{\partial \delta_0 \partial \delta_1} & \frac{\partial^2 l(\delta)}{\partial \delta_1^2} & \dots & \frac{\partial^2 l(\delta)}{\partial \delta_q \partial \delta_1} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^2 l(\delta)}{\partial \delta_0 \partial \delta_q} & \frac{\partial^2 l(\delta)}{\partial \delta_1 \partial \delta_q} & \dots & \frac{\partial^2 l(\delta)}{\partial \delta_q^2} \end{array} \right) \end{aligned}$$

$$= \begin{pmatrix} H_{00} & H_{01} & \cdots & H_{0q} \\ H_{10} & H_{11} & \cdots & H_{1q} \\ \vdots & \vdots & \ddots & \vdots \\ H_{q0} & H_{q1} & \cdots & H_{qq} \end{pmatrix}$$

Dengan menggunakan Metode Newton-Raphson akan dicari taksiran dari δ yang merupakan penyelesaian dari $s(\delta) = 0$. Jika diberikan taksiran awal dari δ , yaitu $\hat{\delta}^{(0)}$, maka dapat diperoleh pendekatan Deret Taylor Orde Pertama dari vektor $s(\delta)$ sekitar $\delta = \hat{\delta}^{(0)}$, yaitu:

$$\begin{aligned} s(\delta) &\approx s(\hat{\delta}^{(0)}) + \left. \frac{\partial s(\delta)}{\partial \delta_0} \right|_{\delta=\hat{\delta}^{(0)}} (\delta_0 - \hat{\delta}_0^{(0)}) + \left. \frac{\partial s(\delta)}{\partial \delta_1} \right|_{\delta=\hat{\delta}^{(0)}} (\delta_1 - \hat{\delta}_1^{(0)}) + \dots + \left. \frac{\partial s(\delta)}{\partial \delta_q} \right|_{\delta=\hat{\delta}^{(0)}} (\delta_q - \hat{\delta}_q^{(0)}) \\ &\approx s(\hat{\delta}^{(0)}) + \begin{pmatrix} \left. \frac{\partial s(\delta)}{\partial \delta_0} \right|_{\delta=\hat{\delta}^{(0)}} & \left. \frac{\partial s(\delta)}{\partial \delta_1} \right|_{\delta=\hat{\delta}^{(0)}} & \cdots & \left. \frac{\partial s(\delta)}{\partial \delta_q} \right|_{\delta=\hat{\delta}^{(0)}} \end{pmatrix} \begin{pmatrix} (\delta_0 - \hat{\delta}_0^{(0)}) \\ (\delta_1 - \hat{\delta}_1^{(0)}) \\ \vdots \\ (\delta_q - \hat{\delta}_q^{(0)}) \end{pmatrix} \\ &\approx s(\hat{\delta}^{(0)}) + \begin{pmatrix} \frac{\partial s_0(\delta)}{\partial \delta_0} & \frac{\partial s_0(\delta)}{\partial \delta_1} & \cdots & \frac{\partial s_0(\delta)}{\partial \delta_q} \\ \frac{\partial s_1(\delta)}{\partial \delta_0} & \frac{\partial s_1(\delta)}{\partial \delta_1} & \cdots & \frac{\partial s_1(\delta)}{\partial \delta_q} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial s_q(\delta)}{\partial \delta_0} & \frac{\partial s_q(\delta)}{\partial \delta_1} & \cdots & \frac{\partial s_q(\delta)}{\partial \delta_q} \end{pmatrix}_{\delta=\hat{\delta}^{(0)}} (\delta - \hat{\delta}^{(0)}) \end{aligned}$$

$$\approx \mathbf{s}(\hat{\boldsymbol{\delta}}^{(0)}) + \begin{pmatrix} \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_0^2} & \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_1 \partial \delta_0} & \dots & \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_q \partial \delta_0} \\ \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_0 \partial \delta_1} & \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_1^2} & \dots & \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_q \partial \delta_1} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_0 \partial \delta_q} & \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_1 \partial \delta_q} & \dots & \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_q^2} \end{pmatrix}_{\boldsymbol{\delta}=\hat{\boldsymbol{\delta}}^{(0)}} (\boldsymbol{\delta} - \hat{\boldsymbol{\delta}}^{(0)})$$

$$\approx \mathbf{s}(\hat{\boldsymbol{\delta}}^{(0)}) + \begin{pmatrix} H_{00}(\boldsymbol{\delta}) & H_{01}(\boldsymbol{\delta}) & \dots & H_{0q}(\boldsymbol{\delta}) \\ H_{10}(\boldsymbol{\delta}) & H_{11}(\boldsymbol{\delta}) & \dots & H_{1q}(\boldsymbol{\delta}) \\ \vdots & \vdots & \ddots & \vdots \\ H_{q0}(\boldsymbol{\delta}) & H_{q1}(\boldsymbol{\delta}) & \dots & H_{qq}(\boldsymbol{\delta}) \end{pmatrix}_{\boldsymbol{\delta}=\hat{\boldsymbol{\delta}}^{(0)}} (\boldsymbol{\delta} - \hat{\boldsymbol{\delta}}^{(0)})$$

atau

$$\mathbf{s}(\boldsymbol{\delta}) \approx \mathbf{s}(\hat{\boldsymbol{\delta}}^{(0)}) + \mathbf{H}(\hat{\boldsymbol{\delta}}^{(0)}) (\boldsymbol{\delta} - \hat{\boldsymbol{\delta}}^{(0)})$$

dengan

$\mathbf{s}(\hat{\boldsymbol{\delta}}^{(0)})$ adalah vektor turunan parsial pertama dari $l(\boldsymbol{\delta})$ dengan elemen

$$s_j(\boldsymbol{\delta}) = \partial l(\boldsymbol{\delta}) / \partial \delta_j \text{ dihitung pada } \boldsymbol{\delta} = \hat{\boldsymbol{\delta}}^{(0)}.$$

$\mathbf{H}(\hat{\boldsymbol{\delta}}^{(0)}) = \left. \frac{\partial \mathbf{s}(\boldsymbol{\delta})}{\partial \boldsymbol{\delta}^T} \right|_{\boldsymbol{\delta}=\hat{\boldsymbol{\delta}}^{(0)}}$ adalah matriks turunan parsial kedua dari $l(\boldsymbol{\delta})$ dengan

elemen $H_{jk}(\boldsymbol{\delta}) = \frac{\partial^2 l(\boldsymbol{\delta})}{\partial \delta_k \partial \delta_j}$ dihitung pada $\boldsymbol{\delta} = \hat{\boldsymbol{\delta}}^{(0)}$.

Oleh karena ingin dicari penyelesaian dari $\mathbf{s}(\boldsymbol{\delta}) = 0$, maka dapat ditulis

$$\begin{aligned} \mathbf{0} &= \mathbf{s}(\delta) \approx \mathbf{s}(\hat{\delta}^{(0)}) + \mathbf{H}(\hat{\delta}^{(0)})(\delta - \hat{\delta}^{(0)}) \\ \mathbf{H}(\hat{\delta}^{(0)}) &(\delta - \hat{\delta}^{(0)}) = -\mathbf{s}(\hat{\delta}^{(0)}). \end{aligned}$$

Jika matriks $\mathbf{H}(\hat{\delta}^{(0)})$ nonsingular maka

$$\begin{aligned} (\mathbf{H}(\hat{\delta}^{(0)}))^{-1} \mathbf{H}(\hat{\delta}^{(0)}) (\delta - \hat{\delta}^{(0)}) &= -(\mathbf{H}(\hat{\delta}^{(0)}))^{-1} \mathbf{s}(\hat{\delta}^{(0)}) \\ (\delta - \hat{\delta}^{(0)}) &= -(\mathbf{H}(\hat{\delta}^{(0)}))^{-1} \mathbf{s}(\hat{\delta}^{(0)}) \\ \delta &= \hat{\delta}^{(0)} - (\mathbf{H}(\hat{\delta}^{(0)}))^{-1} \mathbf{s}(\hat{\delta}^{(0)}). \end{aligned}$$

Nilai δ menjadi taksiran baru yang kemudian dinotasikan dengan $\hat{\delta}^{(1)}$, yaitu:

$$\hat{\delta}^{(1)} = \hat{\delta}^{(0)} - (\mathbf{H}(\hat{\delta}^{(0)}))^{-1} \mathbf{s}(\hat{\delta}^{(0)})$$

Jika $\hat{\delta}^{(1)} \approx \hat{\delta}^{(0)}$ (misalkan $\|\hat{\delta}^{(1)} - \hat{\delta}^{(0)}\| < 10^{-5}$) maka hentikan proses iterasi dan kemudian ambil $\hat{\delta}^{(1)}$ sebagai taksiran δ . Sebaliknya, lanjutkan iterasi.

Jika diberikan $\hat{\delta}^{(1)}$, maka dapat diperoleh pendekatan Deret Taylor Orde Pertama dari vektor $\mathbf{s}(\delta)$ sekitar $\delta = \hat{\delta}^{(1)}$, yaitu:

$$\mathbf{s}(\delta) \approx \mathbf{s}(\hat{\delta}^{(1)}) + \mathbf{H}(\hat{\delta}^{(1)})(\delta - \hat{\delta}^{(1)})$$

dengan

$s(\hat{\delta}^{(1)})$ adalah vektor turunan parsial pertama dari $l(\delta)$ dengan elemen

$s_j(\delta) = \partial l(\delta)/\partial \delta_j$ dihitung pada $\delta = \hat{\delta}^{(1)}$.

$\mathbf{H}(\hat{\delta}^{(1)}) = \left. \frac{\partial s(\delta)}{\partial \delta^T} \right|_{\delta=\hat{\delta}^{(1)}}$ adalah matriks turunan parsial kedua dari $l(\delta)$ dengan

elemen $H_{jk}(\delta) = \frac{\partial^2 l(\delta)}{\partial \delta_k \partial \delta_j}$ dihitung pada $\delta = \hat{\delta}^{(1)}$.

Dengan menyelesaikan $\mathbf{0} = s(\delta) \approx s(\hat{\delta}^{(1)}) + \mathbf{H}(\hat{\delta}^{(1)})(\delta - \hat{\delta}^{(1)})$

diperoleh $\mathbf{H}(\hat{\delta}^{(1)})(\delta - \hat{\delta}^{(1)}) = -s(\hat{\delta}^{(1)})$.

Jika matriks $\mathbf{H}(\hat{\delta}^{(1)})$ nonsingular maka

$$\begin{aligned} (\mathbf{H}(\hat{\delta}^{(1)}))^{-1} \mathbf{H}(\hat{\delta}^{(1)})(\delta - \hat{\delta}^{(1)}) &= -(\mathbf{H}(\hat{\delta}^{(1)}))^{-1} \cdot s(\hat{\delta}^{(1)}) \\ (\delta - \hat{\delta}^{(1)}) &= -(\mathbf{H}(\hat{\delta}^{(1)}))^{-1} \cdot s(\hat{\delta}^{(1)}) \\ \delta &= \hat{\delta}^{(1)} - (\mathbf{H}(\hat{\delta}^{(1)}))^{-1} \cdot s(\hat{\delta}^{(1)}). \end{aligned}$$

Nilai δ menjadi taksiran baru yang kemudian dinotasikan dengan $\hat{\delta}^{(2)}$, yaitu:

$$\hat{\delta}^{(2)} = \hat{\delta}^{(1)} - (\mathbf{H}(\hat{\delta}^{(1)}))^{-1} \cdot s(\hat{\delta}^{(1)})$$

Dengan mengulang langkah yang sama dapat diperoleh taksiran $\hat{\delta}^{(3)}$, $\hat{\delta}^{(4)}$, dan seterusnya. Secara umum dapat ditentukan taksiran dari δ pada iterasi ke- $m+1$, yaitu $\hat{\delta}^{(m+1)}$, secara iteratif menggunakan formula:

$$\hat{\delta}^{(m+1)} = \hat{\delta}^{(m)} - \left(\mathbf{H}(\hat{\delta}^{(m)}) \right)^{-1} \cdot \mathbf{s}(\hat{\delta}^{(m)}); m = 0, 1, \dots$$

dengan

$\mathbf{s}(\hat{\delta}^{(m)})$ adalah vektor turunan parsial pertama dari $l(\delta)$ dengan elemen

$s_j(\delta) = \partial l(\delta)/\partial \delta_j$ dihitung pada $\delta = \hat{\delta}^{(m)}$.

$\mathbf{H}(\hat{\delta}^{(m)})$ adalah matriks turunan parsial kedua dari $l(\delta)$ dengan elemen

$H_{jk}(\delta) = \frac{\partial^2 l(\delta)}{\partial \delta_k \partial \delta_j}$ dihitung pada $\delta = \hat{\delta}^{(m)}$.

Hentikan proses iterasi jika $\hat{\delta}^{(m+1)} \approx \hat{\delta}^{(m)}$ (misalkan $\|\hat{\delta}^{(m+1)} - \hat{\delta}^{(m)}\| < 10^{-5}$), lalu

ambil $\hat{\delta}^{(m+1)}$ sebagai taksiran dari δ .

LAMPIRAN 8

Bukti (2.7.3.b).

(\rightarrow) Diketahui $T(\mathbf{y}) = c + \mathbf{l}^T \mathbf{y}$ dan $E(T(\mathbf{y})) = E(\lambda^T \boldsymbol{\alpha} + \boldsymbol{\omega}^T \mathbf{b})$ maka

$$\begin{aligned}
 E(T(\mathbf{y})) &= E(c + \mathbf{l}^T \mathbf{y}) \\
 &= E(c) + E(\mathbf{l}^T (\mathbf{X}\boldsymbol{\alpha} + \mathbf{Z}\mathbf{b} + \mathbf{e})) \\
 &= c + E(\mathbf{l}^T \mathbf{X}\boldsymbol{\alpha}) + E(\mathbf{l}^T \mathbf{Z}\mathbf{b}) + E(\mathbf{l}^T \mathbf{e}) \\
 &= c + \mathbf{l}^T \mathbf{X}\boldsymbol{\alpha} + \mathbf{l}^T \mathbf{Z}E(\mathbf{b}) + \mathbf{l}^T E(\mathbf{e}) \\
 &= c + \mathbf{l}^T \mathbf{X}\boldsymbol{\alpha} \quad [\text{karena } E(\mathbf{b}) = \mathbf{0} \text{ dan } E(\mathbf{e}) = \mathbf{0}] \tag{L8.1}
 \end{aligned}$$

sedangkan

$$\begin{aligned}
 E(\lambda^T \boldsymbol{\alpha} + \boldsymbol{\omega}^T \mathbf{b}) &= E(\lambda^T \boldsymbol{\alpha}) + E(\boldsymbol{\omega}^T \mathbf{b}) \\
 &= \lambda^T \boldsymbol{\alpha} + \boldsymbol{\omega}^T E(\mathbf{b}) \\
 &= \lambda^T \boldsymbol{\alpha} \quad [\text{karena } E(\mathbf{b}) = \mathbf{0}]. \tag{L8.2}
 \end{aligned}$$

Oleh karena $T(\mathbf{y})$ unbiased terhadap $\lambda^T \boldsymbol{\alpha} + \boldsymbol{\omega}^T \beta$, terlihat bahwa persamaan

(L8.1) akan sama dengan (L8.2) jika $c = 0$ dan $\lambda^T = \mathbf{l}^T \mathbf{X}$.

(\leftarrow) Diketahui bahwa $c = 0$ dan $\lambda^T = \mathbf{l}^T \mathbf{X}$ maka

$$\begin{aligned}
 T(\mathbf{y}) &= c + \mathbf{l}^T \mathbf{y} \\
 &= \mathbf{l}^T (\mathbf{X}\boldsymbol{\alpha} + \mathbf{Z}\mathbf{b} + \mathbf{e}) \\
 &= \mathbf{l}^T \mathbf{X}\boldsymbol{\alpha} + \mathbf{l}^T \mathbf{Z}\mathbf{b} + \mathbf{l}^T \mathbf{e}.
 \end{aligned}$$

Berikut ini akan dibuktikan bahwa $T(\mathbf{y})$ dapat mengestimasi kombinasi linier

$$\lambda^T \alpha + \omega^T \beta$$

$$\begin{aligned}
 E(T(\mathbf{y})) &= E(l^T \mathbf{X}\alpha + l^T \mathbf{Z}\beta + l^T \epsilon) \\
 &= E(l^T \mathbf{X}\alpha) + E(l^T \mathbf{Z}\beta) + E(l^T \epsilon) \\
 &= l^T \mathbf{X}\alpha + l^T \mathbf{Z}E(\beta) + l^T E(\epsilon) \\
 &= l^T \mathbf{X}\alpha \quad [\text{karena } E(\beta) = \mathbf{0} \text{ dan } E(\epsilon) = \mathbf{0}] \\
 &= \lambda^T \alpha \quad [\text{karena } \lambda^T = l^T \mathbf{X}]
 \end{aligned}$$

Jadi, terbukti bahwa $T(\mathbf{y})$ *unbiased* terhadap kombinasi linier $\lambda^T \alpha + \omega^T \beta$.

Selanjutnya, akan dibuktikan bahwa $T(\mathbf{y})$ merupakan penaksir yang linier, yaitu: $T(a\mathbf{y}_1 + b\mathbf{y}_2) = aT(\mathbf{y}_1) + bT(\mathbf{y}_2)$ dengan $T(\mathbf{y}) = l^T \mathbf{y}$.

$$\begin{aligned}
 T(a\mathbf{y}_1 + b\mathbf{y}_2) &= l^T(a\mathbf{y}_1 + b\mathbf{y}_2) \\
 &= l^T a\mathbf{y}_1 + l^T b\mathbf{y}_2 \\
 &= a(l^T \mathbf{y}_1) + b(l^T \mathbf{y}_2) \\
 &= aT(\mathbf{y}_1) + bT(\mathbf{y}_2)
 \end{aligned} \tag{L8.3}$$

Terlihat dari (L8.3) bahwa $T(\mathbf{y})$ merupakan penaksir yang linier terhadap kombinasi linier $\lambda^T \alpha + \omega^T \beta$. Oleh karena $T(\mathbf{y})$ merupakan penaksir yang linier dan *unbiased* terhadap kombinasi linier $\lambda^T \alpha + \omega^T \beta$, maka terbukti bahwa $\lambda^T \alpha + \omega^T \beta$ dapat diestimasi oleh $T(\mathbf{y})$.

LAMPIRAN 9

Bukti (3.1.2).

$$\begin{aligned}
 E(\mathbf{y} | \mathbf{b} = \boldsymbol{\beta}) &= E(\mathbf{X}\boldsymbol{\alpha} + \mathbf{Z}\mathbf{b} + \mathbf{e} | \mathbf{b} = \boldsymbol{\beta}) \\
 &= E(\mathbf{X}\boldsymbol{\alpha} | \mathbf{b} = \boldsymbol{\beta}) + E(\mathbf{Z}\mathbf{b} | \mathbf{b} = \boldsymbol{\beta}) + E(\mathbf{e} | \mathbf{b} = \boldsymbol{\beta}) \\
 &= \mathbf{X}\boldsymbol{\alpha} + \mathbf{Z}\boldsymbol{\beta} \quad [\text{karena } E(\mathbf{e}) = \mathbf{0}]
 \end{aligned}$$

$$\begin{aligned}
 \text{Var}(\mathbf{y} | \mathbf{b} = \boldsymbol{\beta}) &= E((\mathbf{y} - \mathbf{X}\boldsymbol{\alpha} - \mathbf{Z}\mathbf{b})(\mathbf{y} - \mathbf{X}\boldsymbol{\alpha} - \mathbf{Z}\mathbf{b})^T | \mathbf{b} = \boldsymbol{\beta}) \\
 &= E(\mathbf{e}\mathbf{e}^T | \mathbf{b} = \boldsymbol{\beta}) \\
 &= \mathbf{R}.
 \end{aligned}$$

Oleh karena \mathbf{b} dan \mathbf{e} berdistribusi normal, maka $\mathbf{y}|\mathbf{b}$ pun berdistribusi normal atau dari (2.4.b) ditulis sebagai $\mathbf{y}|\mathbf{b} \sim N(\mathbf{X}\boldsymbol{\alpha} + \mathbf{Z}\boldsymbol{\beta}, \mathbf{R})$ sehingga dari (2.4.a) *pdf*-nya adalah sebagai berikut:

$$\begin{aligned}
 g(\mathbf{y} | \boldsymbol{\beta}) &= \frac{1}{(2\pi)^{n/2} |\mathbf{R}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha} - \mathbf{Z}\boldsymbol{\beta})^T \mathbf{R}^{-1} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha} - \mathbf{Z}\boldsymbol{\beta})\right) \\
 &= \frac{1}{(2\pi)^{n/2} |\mathbf{R}|^{1/2}} \exp\left(-\frac{1}{2} \boldsymbol{\varepsilon}^T \mathbf{R}^{-1} \boldsymbol{\varepsilon}\right) \\
 &= g(\boldsymbol{\varepsilon}).
 \end{aligned}$$

LAMPIRAN 10

Bukti (3.1.3).

$$\begin{aligned}
 \ln L(\alpha, \beta; \mathbf{y}) &= -\ln(2\pi)^n (|\mathbf{G}||\mathbf{R}|)^{\frac{1}{2}} - \frac{1}{2} (\boldsymbol{\varepsilon}^T \mathbf{R}^{-1} \boldsymbol{\varepsilon} + \boldsymbol{\beta}^T \mathbf{G}^{-1} \boldsymbol{\beta}) \\
 &= -\ln(2\pi)^n (|\mathbf{G}||\mathbf{R}|)^{\frac{1}{2}} - \frac{1}{2} \left((\mathbf{y} - \mathbf{X}\alpha - \mathbf{Z}\beta)^T \mathbf{R}^{-1} (\mathbf{y} - \mathbf{X}\alpha - \mathbf{Z}\beta) + \boldsymbol{\beta}^T \mathbf{G}^{-1} \boldsymbol{\beta} \right) \\
 &= -\ln(2\pi)^n (|\mathbf{G}||\mathbf{R}|)^{\frac{1}{2}} - \frac{1}{2} \left((\mathbf{y}^T - \boldsymbol{\alpha}^T \mathbf{X}^T - \boldsymbol{\beta}^T \mathbf{Z}^T) \mathbf{R}^{-1} (\mathbf{y} - \mathbf{X}\alpha - \mathbf{Z}\beta) + \boldsymbol{\beta}^T \mathbf{G}^{-1} \boldsymbol{\beta} \right) \\
 &= -\ln(2\pi)^n (|\mathbf{G}||\mathbf{R}|)^{\frac{1}{2}} - \frac{1}{2} (\mathbf{y}^T \mathbf{R}^{-1} \mathbf{y} - \mathbf{y}^T \mathbf{R}^{-1} \mathbf{X}\alpha - \mathbf{y}^T \mathbf{R}^{-1} \mathbf{Z}\beta - \boldsymbol{\alpha}^T \mathbf{X}^T \mathbf{R}^{-1} \mathbf{y} + \boldsymbol{\alpha}^T \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X}\alpha + \boldsymbol{\alpha}^T \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z}\beta \\
 &\quad - \boldsymbol{\beta}^T \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{y} + \boldsymbol{\beta}^T \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X}\alpha + \boldsymbol{\beta}^T \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z}\beta + \boldsymbol{\beta}^T \mathbf{G}^{-1} \boldsymbol{\beta})
 \end{aligned}$$

LAMPIRAN 11

Bukti (3.1.4) dan (3.1.5).

Dengan menggunakan (2.1.3.a), (2.3.2.a), dan (2.3.2.b) didapat

- $\frac{\partial \ln L(\alpha, \beta; y)}{\partial \alpha} = 0$

$$\Leftrightarrow -\left(\mathbf{y}^T \mathbf{R}^{-1} \mathbf{X}\right)^T - \mathbf{X}^T \mathbf{R}^{-1} \mathbf{y} + \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + \left(\tilde{\alpha}^T \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X}\right)^T + \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} + \left(\tilde{\beta}^T \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X}\right)^T = \mathbf{0}$$

$$\Leftrightarrow -\mathbf{X}^T \mathbf{R}^{-1} \mathbf{y} - \mathbf{X}^T \mathbf{R}^{-1} \mathbf{y} + \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} + \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} = \mathbf{0}$$

$$\Leftrightarrow -2\mathbf{X}^T \mathbf{R}^{-1} \mathbf{y} + 2\mathbf{X}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + 2\mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} = \mathbf{0}$$

$$\Leftrightarrow \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} = \mathbf{X}^T \mathbf{R}^{-1} \mathbf{y}$$

- $\frac{\partial \ln L(\alpha, \beta; y)}{\partial \beta} = 0$

$$\Leftrightarrow -\left(\mathbf{y}^T \mathbf{R}^{-1} \mathbf{Z}\right)^T + \left(\tilde{\alpha}^T \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z}\right)^T - \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{y} + \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} + \left(\tilde{\beta}^T \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z}\right)^T + \mathbf{G}^{-1} \tilde{\beta} + \left(\tilde{\beta}^T \mathbf{G}^{-1}\right)^T = \mathbf{0}$$

$$\Leftrightarrow -\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{y} + \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} - \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{y} + \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} + \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} + \mathbf{G}^{-1} \tilde{\beta} + \mathbf{G}^{-1} \tilde{\beta} = \mathbf{0}$$

$$\Leftrightarrow -2\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{y} + 2\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + 2\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} \tilde{\beta} + 2\mathbf{G}^{-1} \tilde{\beta} = \mathbf{0}$$

$$\Leftrightarrow \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} + (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1}) \tilde{\beta} = \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{y}$$

LAMPIRAN 12

Bukti (3.1.6) dan (3.1.7).

Taksiran parameter $\tilde{\alpha}$ dan $\tilde{\beta}$ dengan asumsi **b** dan **e** berdistribusi normal didapat dengan mengeliminasi $\tilde{\beta}$ dari (3.1.4) dan (3.1.5), yaitu mengalikan (3.1.4) dengan **I** dan mengalikan $\mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1}$ dengan (3.1.5), kemudian kedua perkalian tersebut dikurangkan. Hasil eliminasinya adalah sebagai berikut:

$$\begin{aligned} \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} - \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X} \tilde{\alpha} &= \mathbf{X}^T \mathbf{R}^{-1} \mathbf{y} - \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{y} \\ \Leftrightarrow \mathbf{X}^T \left(\mathbf{R}^{-1} - \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \right) \mathbf{X} \tilde{\alpha} &= \mathbf{X}^T \left(\mathbf{R}^{-1} - \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \right) \mathbf{y} \end{aligned}$$

sehingga

$$\tilde{\alpha} = \left(\mathbf{X}^T \left(\mathbf{R}^{-1} - \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \right) \mathbf{X} \right)^{-1} \mathbf{X}^T \left(\mathbf{R}^{-1} - \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \right) \mathbf{y}$$

sedangkan

$$\tilde{\beta} = (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} (\mathbf{y} - \mathbf{X} \tilde{\alpha}).$$

LAMPIRAN 13

Bukti bahwa $\mathbf{W} = \boldsymbol{\Omega}^{-1}$.

$$\begin{aligned}
 \boldsymbol{\Omega}\mathbf{W} &= (\mathbf{R} + \mathbf{ZGZ}^T) \left(\mathbf{R}^{-1} - \mathbf{R}^{-1}\mathbf{Z}(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})^{-1}\mathbf{Z}^T\mathbf{R}^{-1} \right) \\
 &= \mathbf{RR}^{-1} - \mathbf{RR}^{-1}\mathbf{Z}(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})^{-1}\mathbf{Z}^T\mathbf{R}^{-1} + \mathbf{ZGZ}^T\mathbf{R}^{-1} - \mathbf{ZGZ}^T\mathbf{R}^{-1}\mathbf{Z}(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})^{-1}\mathbf{Z}^T\mathbf{R}^{-1} \\
 &= \mathbf{I} - \mathbf{IZ}(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})^{-1}\mathbf{Z}^T\mathbf{R}^{-1} + \mathbf{ZGZ}^T\mathbf{R}^{-1} - \mathbf{ZGZ}^T\mathbf{R}^{-1}\mathbf{Z}(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})^{-1}\mathbf{Z}^T\mathbf{R}^{-1} \\
 &= \mathbf{I} + \mathbf{ZGZ}^T\mathbf{R}^{-1} - \mathbf{Z}(\mathbf{GZ}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{I})(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})^{-1}\mathbf{Z}^T\mathbf{R}^{-1} \\
 &= \mathbf{I} + \mathbf{ZGZ}^T\mathbf{R}^{-1} - \mathbf{ZG}(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})(\mathbf{Z}^T\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1})^{-1}\mathbf{Z}^T\mathbf{R}^{-1} \\
 &= \mathbf{I} + \mathbf{ZGZ}^T\mathbf{R}^{-1} - \mathbf{ZGIZ}^T\mathbf{R}^{-1} \\
 &= \mathbf{I} + \mathbf{ZGZ}^T\mathbf{R}^{-1} - \mathbf{ZGZ}^T\mathbf{R}^{-1} \\
 &= \mathbf{I}
 \end{aligned}$$

LAMPIRAN 14

Bukti bahwa $(\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} = \mathbf{G} \mathbf{Z}^T \boldsymbol{\Omega}^{-1}$.

$$\begin{aligned}
 (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} &= \left((\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} + \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} - \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \right) \mathbf{Z}^T \mathbf{R}^{-1} \\
 &= \left((\mathbf{I} + \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z}) (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} - \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \right) \mathbf{Z}^T \mathbf{R}^{-1} \\
 &= \left((\mathbf{G} \mathbf{G}^{-1} + \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z}) (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} - \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \right) \mathbf{Z}^T \mathbf{R}^{-1} \\
 &= \left(\mathbf{G} (\mathbf{G}^{-1} + \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z}) (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} - \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \right) \mathbf{Z}^T \mathbf{R}^{-1} \\
 &= \left(\mathbf{G} - \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \right) \mathbf{Z}^T \mathbf{R}^{-1} \\
 &= \left(\mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} - \mathbf{G} \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \right) \\
 &= \mathbf{G} \mathbf{Z}^T \left(\mathbf{R}^{-1} - \mathbf{R}^{-1} \mathbf{Z} (\mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1})^{-1} \mathbf{Z}^T \mathbf{R}^{-1} \right) \\
 &= \mathbf{G} \mathbf{Z}^T \boldsymbol{\Omega}^{-1} \quad [\text{karena (3.2.1)}]
 \end{aligned}$$

LAMPIRAN 15

Bukti (3.3.1).

$$\begin{aligned}
 E(\mathbf{y}) &= E(\mathbf{X}\boldsymbol{\alpha} + \mathbf{Z}\mathbf{b} + \mathbf{e}) \\
 &= E(\mathbf{X}\boldsymbol{\alpha}) + E(\mathbf{Z}\mathbf{b}) + E(\mathbf{e}) \\
 &= \mathbf{X}\boldsymbol{\alpha} + \mathbf{Z}E(\mathbf{b}) + E(\mathbf{e}) \\
 &= \mathbf{X}\boldsymbol{\alpha} \quad [\text{karena } E(\mathbf{b}) = \mathbf{0} \text{ dan } E(\mathbf{e}) = \mathbf{0}] \\
 \\
 Var(\mathbf{y}) &= E((\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})(\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T) \\
 &= E((\mathbf{Z}\mathbf{b} + \mathbf{e})(\mathbf{Z}\mathbf{b} + \mathbf{e})^T) \\
 &= E((\mathbf{Z}\mathbf{b} + \mathbf{e})(\mathbf{b}^T \mathbf{Z}^T + \mathbf{e}^T)) \\
 &= E(\mathbf{Z}\mathbf{b}\mathbf{b}^T \mathbf{Z}^T + \mathbf{Z}\mathbf{b}\mathbf{e}^T + \mathbf{e}\mathbf{b}^T \mathbf{Z}^T + \mathbf{e}\mathbf{e}^T) \\
 &= E(\mathbf{Z}\mathbf{b}\mathbf{b}^T \mathbf{Z}^T) + E(\mathbf{Z}\mathbf{b}\mathbf{e}^T) + E(\mathbf{e}\mathbf{b}^T \mathbf{Z}^T) + E(\mathbf{e}\mathbf{e}^T) \\
 &= \mathbf{Z}E(\mathbf{b}\mathbf{b}^T)\mathbf{Z}^T + \mathbf{Z}E(\mathbf{b}\mathbf{e}^T) + E(\mathbf{e}\mathbf{b}^T)\mathbf{Z}^T + \mathbf{R} \\
 &= \mathbf{Z}\mathbf{G}\mathbf{Z}^T + \mathbf{R} \quad [\text{karena } E(\mathbf{b}\mathbf{e}^T) = E(\mathbf{e}\mathbf{b}^T) = \mathbf{0}] \\
 &= \boldsymbol{\Omega} \quad [\text{karena } \boldsymbol{\Omega} = \mathbf{R} + \mathbf{Z}\mathbf{G}\mathbf{Z}^T.]
 \end{aligned}$$

Oleh karena \mathbf{b} dan \mathbf{e} berdistribusi normal, maka \mathbf{y} pun berdistribusi normal atau ditulis sebagai $\mathbf{y} \sim N(\mathbf{X}\boldsymbol{\alpha}, \boldsymbol{\Omega})$ di mana pdf-nya adalah sebagai berikut:

$$f(\mathbf{y}) = \frac{1}{(2\pi)^{n/2} |\boldsymbol{\Omega}(\boldsymbol{\delta})|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T \boldsymbol{\Omega}^{-1}(\boldsymbol{\delta}) (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})\right)$$

LAMPIRAN 16

Bukti (3.3.3).

$$\begin{aligned}
 s_j(\boldsymbol{\alpha}, \boldsymbol{\delta}) &= \frac{\partial \ln L(\boldsymbol{\alpha}, \boldsymbol{\delta}; \mathbf{y})}{\partial \delta_j} = \frac{\partial \left(c - \frac{1}{2} \left(\ln(\|\boldsymbol{\Omega}(\boldsymbol{\delta})\|) + (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T \boldsymbol{\Omega}^{-1}(\boldsymbol{\delta})(\mathbf{y} - \mathbf{X}\boldsymbol{\alpha}) \right) \right)}{\partial \delta_j} \\
 &= \frac{\partial c}{\partial \delta_j} - \frac{1}{2} \frac{\partial \ln(\|\boldsymbol{\Omega}(\boldsymbol{\delta})\|)}{\partial \delta_j} - \frac{1}{2} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T \frac{\partial \boldsymbol{\Omega}^{-1}(\boldsymbol{\delta})}{\partial \delta_j} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha}) \\
 &= -\frac{1}{2} \text{tr} \left(\boldsymbol{\Omega}^{-1}(\boldsymbol{\delta}) \frac{\partial \boldsymbol{\Omega}(\boldsymbol{\delta})}{\partial \delta_j} \right) + \frac{1}{2} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T \left(\boldsymbol{\Omega}^{-1}(\boldsymbol{\delta}) \frac{\partial \boldsymbol{\Omega}(\boldsymbol{\delta})}{\partial \delta_j} \boldsymbol{\Omega}^{-1}(\boldsymbol{\delta}) \right) (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})
 \end{aligned}$$

LAMPIRAN 17

Bukti (3.3.5).

Akan dicari $\Upsilon(\delta)$ sehingga *Scoring Algorithm* dapat digunakan. Misalkan

$\Upsilon_{jk}(\delta)$ adalah elemen ke- jk dari $\Upsilon(\delta)$, yaitu:

$$\Upsilon_{jk}(\delta) = -E\left(\frac{\partial^2 \ln L(\alpha, \delta; \mathbf{y})}{\partial \delta_j \partial \delta_k}\right)$$

sedangkan $\Omega_{(j)} = \frac{\partial \Omega(\delta)}{\partial \delta_j}$ dan $\Omega_{(j,k)} = \frac{\partial \Omega(\delta)}{\partial \delta_k \partial \delta_j}$.

$$\begin{aligned} \Upsilon_{jk}(\delta) &= -E\left(\frac{\partial^2 \left(c - \frac{1}{2} \left(\ln(|\Omega|) + (\mathbf{y} - \mathbf{X}\alpha)^T \Omega^{-1} (\mathbf{y} - \mathbf{X}\alpha)\right)\right)}{\partial \delta_k \partial \delta_j}\right) \\ &= -E\left(\frac{\partial^2 c}{\partial \delta_k \partial \delta_j} - \frac{1}{2} \frac{\partial^2 \ln(|\Omega|)}{\partial \delta_k \partial \delta_j} - \frac{1}{2} (\mathbf{y} - \mathbf{X}\alpha)^T \frac{\partial^2 \Omega^{-1}}{\partial \delta_k \partial \delta_j} (\mathbf{y} - \mathbf{X}\alpha)\right) \\ &= -E\left(-\frac{1}{2} \left(tr\left(\Omega^{-1} \frac{\partial^2 \Omega}{\partial \delta_k \partial \delta_j}\right) - tr\left(\Omega^{-1} \frac{\partial \Omega}{\partial \delta_k} \Omega^{-1} \frac{\partial \Omega}{\partial \delta_j}\right)\right) - \frac{1}{2} (\mathbf{y} - \mathbf{X}\alpha)^T \right. \\ &\quad \left. \left(-\Omega^{-1} \frac{\partial^2 \Omega}{\partial \delta_k \partial \delta_j} \Omega^{-1} + \Omega^{-1} \frac{\partial \Omega}{\partial \delta_k} \Omega^{-1} \frac{\partial \Omega}{\partial \delta_j} \Omega^{-1} + \Omega^{-1} \frac{\partial \Omega}{\partial \delta_j} \Omega^{-1} \frac{\partial \Omega}{\partial \delta_k} \Omega^{-1}\right) (\mathbf{y} - \mathbf{X}\alpha)\right) \\ &= -E\left(-\frac{1}{2} \left(tr\left(\Omega^{-1} \Omega_{(j,k)}\right) - tr\left(\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)}\right)\right) - \frac{1}{2} (\mathbf{y} - \mathbf{X}\alpha)^T \right. \\ &\quad \left. \left(-\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} + \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} + \Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1}\right) (\mathbf{y} - \mathbf{X}\alpha)\right) \\ &= -E\left(-\frac{1}{2} \left(tr\left(\Omega^{-1} \Omega_{(j,k)}\right) - tr\left(\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)}\right)\right) - \frac{1}{2} tr\left((\mathbf{y} - \mathbf{X}\alpha)^T \begin{pmatrix} -\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1} \end{pmatrix} (\mathbf{y} - \mathbf{X}\alpha)\right)\right) \end{aligned}$$

$$\begin{aligned}
&= -E \left(-\frac{1}{2} \left(\text{tr}(\Omega^{-1} \Omega_{(j,k)}) - \text{tr}(\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)}) \right) - \frac{1}{2} \text{tr} \left(\begin{pmatrix} -\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1} \end{pmatrix} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha}) (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T \right) \right) \\
&= -E \left(-\frac{1}{2} \left(\text{tr}(\Omega^{-1} \Omega_{(j,k)}) - \text{tr}(\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)}) + \text{tr} \left(\begin{pmatrix} -\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1} \end{pmatrix} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha}) (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T \right) \right) \right) \\
&= -E \left(-\frac{1}{2} \left(\text{tr} \left(\Omega^{-1} \Omega_{(j,k)} - \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} + \begin{pmatrix} -\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1} \end{pmatrix} (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha}) (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T \right) \right) \right) \\
&= \frac{1}{2} \text{tr} \left(\Omega^{-1} \Omega_{(j,k)} - \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} + \begin{pmatrix} -\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} + \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} \\ +\Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1} \end{pmatrix} E((\mathbf{y} - \mathbf{X}\boldsymbol{\alpha}) (\mathbf{y} - \mathbf{X}\boldsymbol{\alpha})^T) \right) \\
&= \frac{1}{2} \text{tr} \left(\Omega^{-1} \Omega_{(j,k)} - \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} + (-\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} + \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} + \Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1}) \Omega \right) \\
&= \frac{1}{2} \text{tr} \left(\Omega^{-1} \Omega_{(j,k)} - \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} + (-\Omega^{-1} \Omega_{(j,k)} \Omega^{-1} \Omega + \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} \Omega^{-1} + \Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega) \right) \\
&= \frac{1}{2} \text{tr} \left(\Omega^{-1} \Omega_{(j,k)} - \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} - \Omega^{-1} \Omega_{(j,k)} + \Omega^{-1} \Omega_{(k)} \Omega^{-1} \Omega_{(j)} + \Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \right) \\
&= \frac{1}{2} \text{tr} \left(\Omega^{-1} \Omega_{(j)} \Omega^{-1} \Omega_{(k)} \right).
\end{aligned}$$

LAMPIRAN 18

Bukti bahwa $\hat{\tau}(\hat{\delta}(y), y)$ Tetap *Unbiased* untuk Mengestimasi $\tau = \lambda^T \alpha + \omega^T \beta$.

Dengan menggunakan (2.7.3.c) dan (2.7.3.d) maka didapatkan

$$\begin{aligned}
 \lambda^T \hat{\alpha}(\delta, y) - \lambda^T \alpha &= \lambda^T (\hat{\alpha}(\delta, y) - \alpha) \\
 &= \lambda^T \left((\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) y - (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X} \alpha \right) \\
 &= \lambda^T (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) (y - \mathbf{X} \alpha) \\
 &= \lambda^T (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) (Zb + e)
 \end{aligned} \tag{L18.1}$$

$$\begin{aligned}
 \hat{\beta}(\delta, y) &= \mathbf{G}(\delta) Z^T \Omega^{-1}(\delta) (y - \mathbf{X} \hat{\alpha}(\delta, y)) \\
 &= \mathbf{G}(\delta) Z^T \Omega^{-1}(\delta) \left(y - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) y \right) \\
 &= \mathbf{G}(\delta) Z^T \Omega^{-1}(\delta) \left(I - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) \right) y \\
 &= \mathbf{G}(\delta) Z^T \Omega^{-1}(\delta) \left(I - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) \right) (X\alpha + Zb + e) \\
 &= \mathbf{G}(\delta) Z^T \Omega^{-1}(\delta) \left(\begin{array}{l} \mathbf{X}\alpha - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) X\alpha + Zb + e \\ - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) (Zb + e) \end{array} \right) \\
 &= \mathbf{G}(\delta) Z^T \Omega^{-1}(\delta) \left(\mathbf{X}\alpha - \mathbf{X}\alpha + \left(I - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) \right) (Zb + e) \right) \\
 &= \mathbf{G}(\delta) Z^T \Omega^{-1}(\delta) \left(I - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta) \right) (Zb + e).
 \end{aligned} \tag{L18.2}$$

Dengan memisalkan $\mathbf{P}(\delta) = I - \mathbf{X} (\mathbf{X}^T \Omega^{-1}(\delta) \mathbf{X})^{-1} \mathbf{X}^T \Omega^{-1}(\delta)$ maka persamaan (L18.2) dapat ditulis menjadi

$$\hat{\beta}(\delta, \mathbf{y}) = \mathbf{G}(\delta) \mathbf{Z}^T \Omega^{-1}(\delta) \mathbf{P}(\delta) (\mathbf{Z}\mathbf{b} + \mathbf{e}). \quad (\text{L18.3})$$

Berdasarkan (L18.1) dan (L18.3) maka

$$\begin{aligned} \hat{\tau}(\hat{\delta}(\mathbf{y}), \mathbf{y}) - t &= \lambda^T \hat{\alpha}(\hat{\delta}(\mathbf{y}), \mathbf{y}) + \omega^T \hat{\beta}(\hat{\delta}(\mathbf{y}), \mathbf{y}) - \lambda^T \alpha - \omega^T \mathbf{b} \\ &= \lambda^T \left(\mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{y})) \mathbf{X} \right)^{-1} \mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{y})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) \\ &\quad + \omega^T \mathbf{G}(\hat{\delta}(\mathbf{y})) \mathbf{Z}^T \Omega^{-1}(\hat{\delta}(\mathbf{y})) \mathbf{P}(\hat{\delta}(\mathbf{y})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) - \omega^T \mathbf{b}. \end{aligned}$$

Dengan menggunakan (3.4.2) dan memisalkan

$$\begin{aligned} \phi(\mathbf{b}, \mathbf{e}) &= \lambda^T \left(\mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{X} \right)^{-1} \mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) \\ &\quad + \omega^T \mathbf{G}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{Z}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{P}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) \end{aligned}$$

maka $\hat{\tau}(\hat{\delta}(\mathbf{y}), \mathbf{y}) - t = \phi(\mathbf{b}, \mathbf{e}) - \omega^T \mathbf{b}$. Berdasarkan (3.4.2) juga didapat

$$\begin{aligned} \phi(-\mathbf{b}, -\mathbf{e}) &= \lambda^T \left(\mathbf{X}^T \Omega^{-1}(\hat{\delta}(-(\mathbf{Z}\mathbf{b} + \mathbf{e}))) \mathbf{X} \right)^{-1} \mathbf{X}^T \Omega^{-1}(\hat{\delta}(-(\mathbf{Z}\mathbf{b} + \mathbf{e}))) (-(\mathbf{Z}\mathbf{b} + \mathbf{e})) \\ &\quad + \omega^T \mathbf{G}(\hat{\delta}(-(\mathbf{Z}\mathbf{b} + \mathbf{e}))) \mathbf{Z}^T \Omega^{-1}(\hat{\delta}(-(\mathbf{Z}\mathbf{b} + \mathbf{e}))) \mathbf{P}(\hat{\delta}(-(\mathbf{Z}\mathbf{b} + \mathbf{e}))) (-(\mathbf{Z}\mathbf{b} + \mathbf{e})) \\ &= -\lambda^T \left(\mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{X} \right)^{-1} \mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) \\ &\quad - \omega^T \mathbf{G}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{Z}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{P}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) \\ &= -\left(\lambda^T \left(\mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{X} \right)^{-1} \mathbf{X}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) \right. \\ &\quad \left. + \omega^T \mathbf{G}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{Z}^T \Omega^{-1}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) \mathbf{P}(\hat{\delta}(\mathbf{Z}\mathbf{b} + \mathbf{e})) (\mathbf{Z}\mathbf{b} + \mathbf{e}) \right) \\ &= -\phi(\mathbf{b}, \mathbf{e}) \end{aligned}$$

sehingga dari (2.8.1.b) diketahui bahwa $\phi(\mathbf{b}, \mathbf{e})$ merupakan fungsi ganjil dari \mathbf{b} dan \mathbf{e} . Dengan asumsi bahwa \mathbf{b} dan \mathbf{e} berdistribusi normal (simetris) dengan *mean* nol, maka berdasarkan lemma 3.1 $\phi(\mathbf{b}, \mathbf{e})$ berdistribusi simetris di sekitar nol. Dengan demikian $E(\phi(\mathbf{b}, \mathbf{e})) = 0$ dan

$$\begin{aligned} E(\hat{\tau}(\hat{\delta}(\mathbf{y}), \mathbf{y}) - t) &= E(\phi(\mathbf{b}, \mathbf{e}) - \boldsymbol{\omega}^T \mathbf{b}) \\ &= E(\phi(\mathbf{b}, \mathbf{e})) - E(\boldsymbol{\omega}^T \mathbf{b}) \\ &= 0 - \boldsymbol{\omega}^T E(\mathbf{b}) \\ &= 0. \end{aligned}$$

LAMPIRAN 19

Data Ilustrasi Penggunaan Metode EBLUP pada *General Linear Mixed Model*
serta Bentuk Matriks **X** dan **Z**.

No.	Location	Champaign	Sales(\$)
1	1	3	267.80
2	1	2	248.84
3	1	2	247.89
4	1	2	251.21
5	1	1	276.48
6	2	1	229.77
7	2	1	231.92
8	2	1	232.52
9	2	3	219.36
10	2	3	212.42
11	3	3	181.26
12	3	1	201.34
13	3	2	143.37
14	3	2	144.46
15	3	3	186.87

No.	Location	Champaign	Sales(\$)
16	4	1	225.26
17	4	1	218.27
18	4	2	181.10
19	4	3	228.91
20	4	3	208.66
21	5	2	198.30
22	5	3	232.66
23	5	2	207.18
24	5	1	237.22
25	5	1	245.47
26	6	2	178.58
27	6	2	170.05
28	6	1	208.67
29	6	2	169.16
30	6	3	185.52

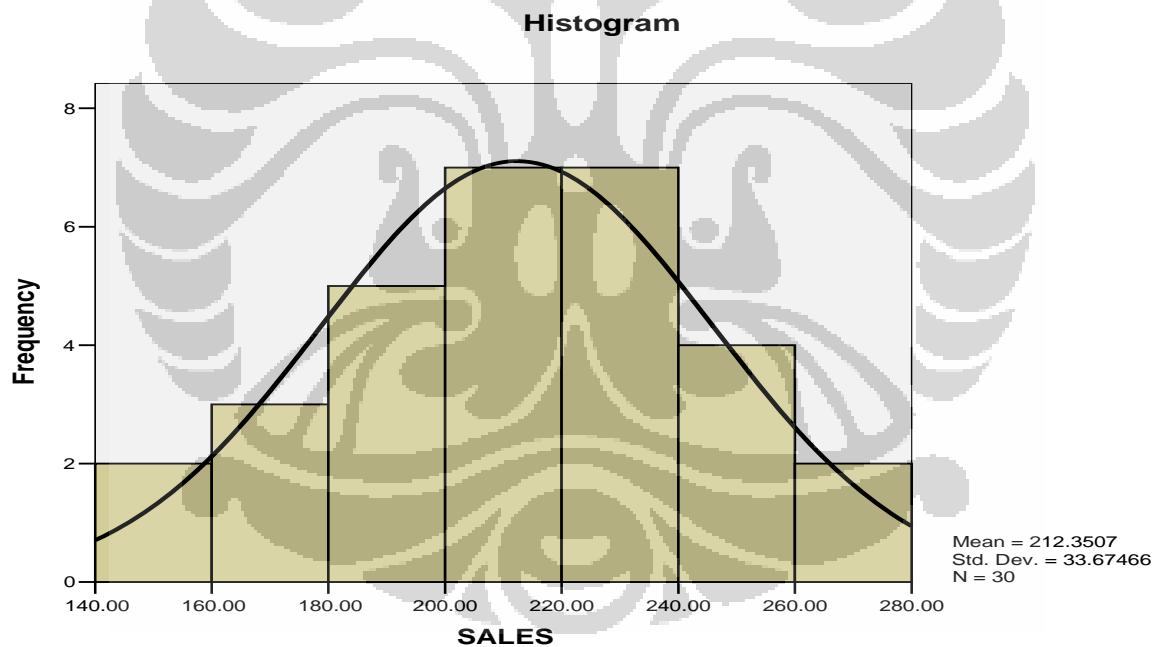
$$\begin{aligned}
 \mathbf{X} = & \begin{pmatrix} 0 & 0 & 1 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 1 & 0 & 0 \\ 0 & 0 & 1 \end{pmatrix} \\
 \mathbf{Z} = & \begin{pmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix}
 \end{aligned}$$

LAMPIRAN 20**Pengujian Kenormalan.****Tests of Normality**

	Kolmogorov-Smirnov ^a			Shapiro-Wilk		
	Statistic	df	Sig.	Statistic	df	Sig.
SALES	.089	30	.200*	.978	30	.778

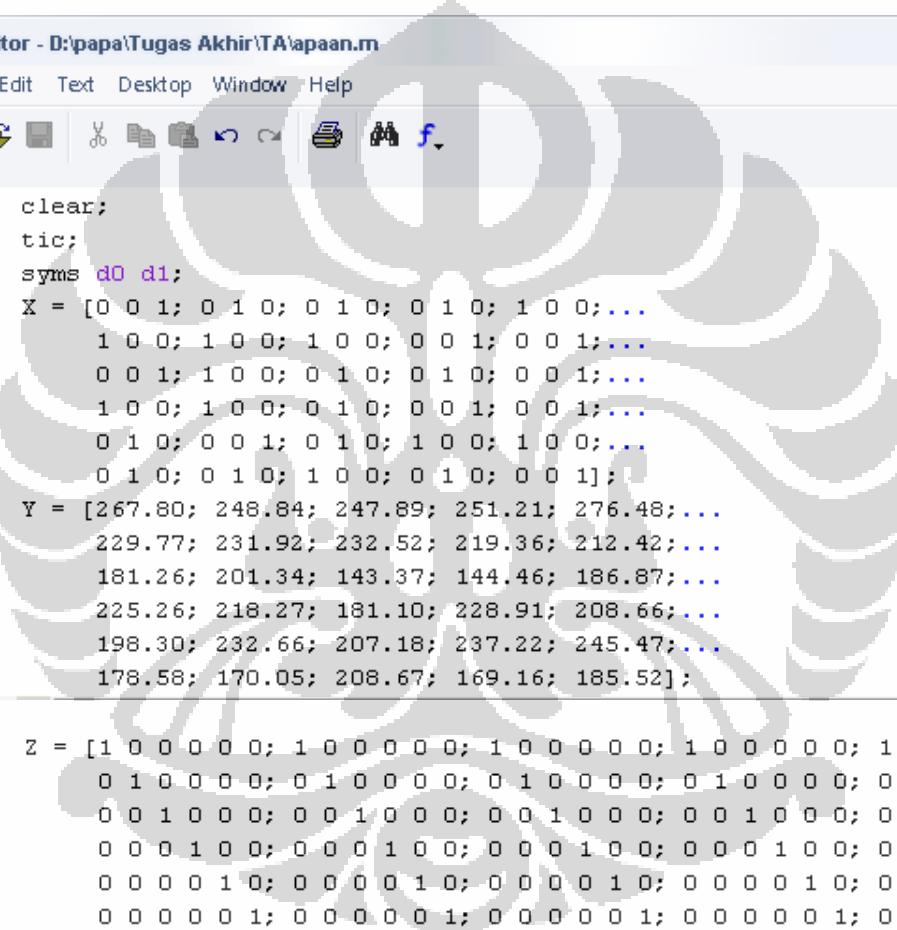
*. This is a lower bound of the true significance.

a. Lilliefors Significance Correction



LAMPIRAN 21

Penaksiran Parameter pada *General Linear Mixed Model* Menggunakan
Software MATLAB 7.0.1.



```
Editor - D:\papa\Tugas Akhir\TA\apaan.m
File Edit Text Desktop Window Help
D F S D F M f
1 clear;
2 tic;
3 syms d0 d1;
4 X = [0 0 1; 0 1 0; 0 1 0; 0 1 0; 1 0 0;...
5     1 0 0; 1 0 0; 1 0 0; 0 0 1; 0 0 1;...
6     0 0 1; 1 0 0; 0 1 0; 0 1 0; 0 0 1;...
7     1 0 0; 1 0 0; 0 1 0; 0 0 1; 0 0 1;...
8     0 1 0; 0 0 1; 0 1 0; 1 0 0; 1 0 0;...
9     0 1 0; 0 1 0; 1 0 0; 0 1 0; 0 0 1];
10 Y = [267.80; 248.84; 247.89; 251.21; 276.48;...
11     229.77; 231.92; 232.52; 219.36; 212.42;...
12     181.26; 201.34; 143.37; 144.46; 186.87;...
13     225.26; 218.27; 181.10; 228.91; 208.66;...
14     198.30; 232.66; 207.18; 237.22; 245.47;...
15     178.58; 170.05; 208.67; 169.16; 185.52];
16 Z = [1 0 0 0 0 0; 1 0 0 0 0 0; 1 0 0 0 0 0; 1 0 0 0 0 0; 1 0 0 0 0 0;...
17     0 1 0 0 0 0; 0 1 0 0 0 0; 0 1 0 0 0 0; 0 1 0 0 0 0; 0 1 0 0 0 0;...
18     0 0 1 0 0 0; 0 0 1 0 0 0; 0 0 1 0 0 0; 0 0 1 0 0 0; 0 0 1 0 0 0;...
19     0 0 0 1 0 0; 0 0 0 1 0 0; 0 0 0 1 0 0; 0 0 0 1 0 0; 0 0 0 1 0 0;...
20     0 0 0 0 1 0; 0 0 0 0 1 0; 0 0 0 0 1 0; 0 0 0 0 1 0; 0 0 0 0 1 0;...
21     0 0 0 0 0 1; 0 0 0 0 0 1; 0 0 0 0 0 1; 0 0 0 0 0 1; 0 0 0 0 0 1];
22 delta = [0.01; 0.01];
23 epsilon = [0.01; 0.01];
24 err = [1; 1];
```

```

26 R = d0*eye(30);
27 G = d1*eye(6);
28 O = R + Z*G*Z';
29 a_cap = inv(X'*inv(O)*X)*(X'*inv(O)*Y);
30 diffy(1,:,:)=diff(O,d0);
31 diffy(2,:,:)=diff(O,d1);
32
33 for j=1:2
34     p = diffy(j,:,:);
35     p = reshape(p, 30, 30);
36     S(j) = -1/2*trace(inv(O)*p) + 1/2*(Y - X*a_cap)'*...
37         (inv(O)*p*inv(O))*(Y-X*a_cap);
38     for k=1:2
39         q = diffy(k,:,:);
40         q = reshape(q, 30, 30);
41         L(j,k) = 1/2*trace((inv(O)*p) * (inv(O)*q));
42     end
43 end

45 while norm(err,inf) >= epsilon
46     nS = subs(S, (d0, d1), (delta(1), delta(2)));
47     nL = subs(L, (d0, d1), (delta(1), delta(2)));
48
49     del = delta;
50     delta = del + inv(nL)*nS';
51
52     err = abs(delta - del);
53 end

55 R = delta(1)*eye(30);
56 G = delta(2)*eye(6);
57 O = R + Z*G*Z';
58 a_cap = inv(X'*inv(O)*X)*(X'*inv(O)*Y)
59 b_cap = G*Z'*inv(O)*(Y - X*a_cap)
60 y_cap = X*a_cap + Z*b_cap
61 Y = y_cap
62 t = toc;
63 fprintf('Waktu: %f', t);

```

taksiranmama.m x MAMA.m x coba.m x apean.m x

script Ln 1 Col 1 OVR ...

LAMPIRAN 22

Tabel Perbandingan antara y dan \hat{y} .

Sales dari Data	Sales dari Taksiran
267.80	271.979
248.84	244.5921
247.89	244.5921
251.21	244.5921
276.48	283.512
229.77	229.8142
231.92	229.8142
232.52	229.8142
219.36	218.2811
212.42	218.2811
181.26	180.5283
201.34	192.0614
143.37	153.1415
144.46	153.1415
186.87	180.5283
225.26	224.8945
218.27	224.8945
181.10	185.9746
228.91	213.3615

208.66	213.3615
198.30	202.9899
232.66	230.3768
207.18	202.9899
237.22	241.9098
245.47	241.9098
178.58	169.3754
170.05	169.3754
208.67	208.2953
169.16	169.3754
185.52	196.7623

LAMPIRAN 23Tabel *Error.*

<i>Error</i>	<i>Error</i>
-4.179	0.3655
4.2479	-6.6245
3.2979	-4.8746
6.6179	15.5485
-7.032	-4.7015
-0.0442	-4.6899
2.1058	2.2832
2.7058	4.1901
1.0789	-4.6898
-5.8611	3.5602
0.7317	9.2046
9.2786	0.6746
-9.7715	0.3747
-8.6815	-0.2154
6.3417	-11.2423